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Moment of Random Variables

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Moment of Random Variables :

Def : - Let a r.v. for c.r.v let $K \in I^+$ Then $E (X^k)$ is called (the K^{th} moment of X) or (the moment of the order K of X)

When $K = 1 \rightarrow E (X^1) = 1^{st}$ moment of $X = M$ $E (X^2) = 2^{nd}$ moment of X

Note : $E(X^K)$ exists if $E (|X|^k) < \infty$

Theorem "1": - If $E (X^k)$ exists then $E (X^j)$

exists , $j < K$ and $j, k \in I^+$

Proof : case "1" If (X) is c.r.v with p.d.f $f (X)$ since $E (X^k)$ exists $\rightarrow \therefore E (|X|^k) < \infty$

T.P $E(X^k)$ exists , T.P $E (|X|^k) < \infty$

$$E (|X|^j) = \int_{-\infty}^{\infty} |X|^j f(x) dx .$$

$$E (|X|^j) = \int_{|X| \leq 1} |X|^j f(X) dx + \int_{|X| > 1} |X|^j f(x) dx$$

2^{nd} central moment of R.V. X is equal to $V(X)$

Ex : - let X be a r.v.s.t. $E (X) = 1$, $E(X^2) = 2$ and $E(X^3) = 5$, Find the 3^{rd} central moment of X

$$\begin{aligned} \text{sol : - } E [(X - M^3)] &= E\{X^3 - 3MX^2 + 3M^2X - M^3\} \\ &= E(X^3) - 3ME(X^2) + 3M^2E(X) - M^3 \\ &= 5 - 3.1.2 + 3.1.1 - 1 = 1 \end{aligned}$$

Moment Generation Function (M. g. f.) .

Def : A moment of generation function(*M. g. f*) of a random variable x is a function that determines all moments of X , denoted by $M_x(t)$, suppose that $t \in (-h, h)$, $h > 0$.

If $E [e^{tx}]$, existis $\forall t \in (-h, h)$ then $M_x(t) = E [e^{tx}]$, $-h < t < h$

There are two case of $M_x (t)$

Case (1)

If X is a discrete random variable from a p.m.f , $f (x)$

$$M_x(t) = E[e^{tx}] = \sum_{\forall x} e^{tx} f(x)$$

Case (2)

If x is a continuous random variable have a p.d.f , $f (x)$

$$M_x(t) = E[t^{tx}] = \int_0^{\infty} e^{tx} f(x) dx$$

Ex : Given a p.d.f , $f (x) = \begin{cases} e^{-x} & \text{for } x > 0 \\ 0 & \text{o.w} \end{cases}$

Find $M_x(t)$

$$\text{sol : } M_x(t) = E (e^{tx}) = \int_0^{\infty} e^{tx} \cdot e^{-x} dx = \int_0^{\infty} e^{-(1-t)x} dx$$

This integration exist only when $(1 - t) > 0$, its mean $t < 1$.

$$M_x(t) = \frac{-1}{1-t} e^{-[1-t]x} \Big|_0^{\infty} = \frac{-1}{1-t} [e^{-\infty} - e^0] = \frac{1}{1-t}$$

$$M_x(t) = \frac{1}{1-t} \quad \text{for } t < 1.$$

Theorem " 2": Let x be a random variable have a M.g.f., then

$$M_x(t = 0) = 1, M'_x(t = 0) = E(x), M''_x(t = 0) = E(x^2),$$

$$M'''_x(t = 0) = E(x^3), \dots M_x^h(t = 0) = E(x^h)$$

proof:

$M_x(t) = E(e^{tx})$, by Macclaurin series

$$e^x = 1 + x + \frac{x^2}{2!} + \frac{x^3}{3!} + \dots + \frac{x^k}{k!} + \dots$$

$$\therefore e^{tx} = 1 + tx + \frac{t^2x^2}{2!} + \frac{t^3x^3}{3!} + \dots + \frac{t^kx^k}{k!} + \dots$$

$$M_x(t) = E\left[1 + tx + \frac{t^2x^2}{2!} + \frac{t^3x^3}{3!} + \dots + \frac{t^kx^k}{k!} + \dots\right]$$

$$M_x(t) = 1 + tEx + \frac{t^2E(x^2)}{2!} + \frac{t^3E(x^3)}{3!} + \dots + \frac{t^kE(x^k)}{k!} + \dots]$$

$$M_x(t = 0) = M_x(0) = 1$$

$$\begin{aligned} M'_x(t) &= \frac{dM_x(t)}{dt} \\ &= E(x) + \frac{2t}{2!} E(x^2) + \frac{3t^2}{3!} E(x^3) + \frac{kt^{k-1}}{k!} E(x^k) + \dots \end{aligned}$$

$$M'_x(t = 0) = E(x)$$

$$M''_x(t = 0) = E(x^2) + \frac{6t}{3!} E(x^3) + \dots + \frac{k(k-1)t^{k-2}}{k!} E(x^k) + \dots$$

$$M''_x(t = 0) = E(x^2)$$

similarly we can find $E(x^3), E(x^4), \dots, E(x^k)$

Note:

$$M_x(t) = M_x(0) + t M'_x(0) + \frac{t^2}{2!} M''_x(0) + \dots + \frac{t^k}{k!} M_x^{(k)}(t=0)$$

This series is called *M. g. f. by Macclaurin series.*

EX.: Given $M_x(t) = \frac{1}{1-2t}$, $t < \frac{1}{2}$, find $E(x)$ and $V(x)$

sol.

$$M_x(t) = (1 - 2t)^{-1}, t < \frac{1}{2}$$

$$M'_x(t) = -(1 - 2t)^{-2} \cdot (-2) = 2(1 - 2t)^{-2}$$

$$E(x) = M'_x(t=0) = 2(1 - 0)^{-2} = 2$$

$$V(x) = E(x^2) - [E(x)]^2$$

$$M''_x(t) = 8(1 - 2t)^{-3}$$

$$E(x^2) = M''_x(0) = 8(1 - 0)^{-3} = 8$$

$$\therefore V(x) = 8 - (2)^2 = 4 \geq 0$$

Theorem "3":

Let x be a random variable have a M. g. f. $M_x(t)$, If $y = ax + b$,

$$a, b \in R, \text{ Then } M_x(t) = e^{bt} M_x(at)$$

proof:

$$y = ax + b$$

$$M_x(t) = E(e^{ty}) = E(e^{t(ax+b)})$$

$$= E[e^{(at)x} \cdot e^{bt}] = e^{bt} E[e^{(at)x}]$$

$$= e^{bt} M_x(at)$$

$$\text{since } M_x(t) = E(e^{tx})$$

$$\rightarrow M_x(at) = E[e^{(at)x}]$$